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MELBOURNE, VICTORIA

STRUCTURES REPORT 381

FATIGUE LIFE VARIABILITY IN ALUMINIUM ALLOY AIRCRAFT STRUCTURES

by

G. S. JOST and S. P. COSTOLLOE



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G. S. JOST and S. P. COSTOLLOE

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SUMMARY

A survey of variability in the fatigue lives of aluminium alloy aircraft structures tested under gust and manoeuvre loadings using programmed and random sequences has shown that scatter associated with gust loading is significantly higher than that for manoeuvre loading. By contrast, there appears to be no systematic effect of loading sequence.

The data have been treated both as lognormal and Weibull distributed

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DOCUMENT CONTROL DATA SHEET

. Document Numbers (a) AR Number: AR-001-788 (b) Document Series and Number:	2. Security Classification (a) Complete document:							
AR-001-788 (b) Document Series and Number:								
	Unclassified							
	(b) Title in isolation:							
Structures Report 381	Unclassified							
(c) Report Number: ARL-Struc-Report-381	(c) Summary in isolation: Unclassified							
B. Title: FATIGUE LIFE VARIABILITY STRUCTURES	IN ALUMINIUM ALLOY AIRCRAFT							
Personal Author(s):	5. Document Date:							
G. S. Jost S. P. Costolloe	January, 1980							
S. F. Costolioe	6. Type of Report and Period Covered:							
Corporate Author(s): Aeronautical Research Laboratories	8. Reference Numbers (a) Task: (b) Sponsoring Agency:							
7. Cost Code: 27 7030	(b) Sponsoring / Iguney.							
. Imprint: Aeronautical Research Laboratories, Melbourne	11. Computer Program(s) (Title(s) and language(s)): Not applicable							
Release Limitations (of the document) Approved for public release								
2-0. Overseas: N.O. P.R. 1	B C D E							
Announcement Limitations (of the inform No limitation	ation on this page):							
. Descriptors:	15. Cosati Codes:							
Fatigue life Aircraft	structures 1106							
Aluminium alloys Gust loa	is 1113							
ABS	RACT							

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NOTATION

f	function
i	failure number, $1 \le i \le n$
k	sample number, $1 \le k \le m$
n	number in the sample
m	number of sample
s	sample standard deviation of log life
x	fatigue life
y	$\log x$
α	dispersion parameter of Weibull distribution
β	location parameter of Weibull distribution
Γ	gamma function
ν_k	degrees of freedom of k^{th} sample = $n_k - 1$
Σ	sum
σ	standard deviation of log normal distribution
μ	mean of log normal distribution
â	pooled sample estimate of σ
â	pooled sample estimate of α .

1. INTRODUCTION

Since an earlier report on this subject was issued in 1971¹, additional data on fatigue life variability in aircraft structures have been published. In Reference 1 data on scatter arising from full scale fatigue tests on many aluminium alloy aircraft structures were analysed on the basis of the log normal distribution. Data pooling permitted the establishment of typical values for given types of loading and/or sequence. It was considered worthwhile repeating the exercise to include the additional data, and to treat the total body of data also in terms of the two parameter Weibull distribution.

2. DATA

The opportunity has been taken to make some improvements to the earlier report. First, neither the constant amplitude data nor the notched specimen data given there are included here. The relevance of such data to service sequence loading and to real structures, respectively, is now recognised as rather remote, and the present adequate and far more appropriately based data make their use unnecessary.

The variable amplitude data of the original report were classified into symmetric and asymmetric loadings, and those which included ground to air cycles. These categories have been retained but the first two have been retitled gust loading and manoeuvre loading respectively. Loadings in the former category are symmetrical about the 1 g level and are typical of civil or transport aircraft, whilst manoeuvre loadings are characterised by a marked asymmetry typical of, for example, fighter aircraft. All of the additional published data fall into the manoeuvre loading category which now contains over 160 individual test results.

Finally, the source data used for subsequent analysis were not included in the earlier report. That omission is rectified here in Tables 1, 2 and 3 where the details of the data from gust loading, manoeuvre loading and from loadings which included ground to air cycles are given.

it is to be noted that the data considered here have all been generated prior to the general introduction of closed loop servo (or load feedback) fatigue testing equipment with its inherent precision and long term stability. The indications from data obtained on specimens and components under more representative testing and/or more modern testing equipment are that the scatter associated with such tests results does, if anything, tend to decrease. The present scatter estimates may therefore be considered as representing upper variability bounds.

3. ANALYSIS

The data have been fitted to the log normal and Weibull distributions. Defining $y = \log x$, where x represents fatigue life in cycles or hours, the normal distribution of y is given by

$$f(y) = (\sigma \sqrt{2\pi})^{-1} \exp \left[-\frac{1}{2}((y-\mu)/\sigma)^2\right]$$

where μ and σ^2 are the mean and variance respectively, their estimators from a sample of size n being given by

$$\bar{y} = \sum_{i=1}^{n} y_i / n$$

$$s^2 = \sum_{i=1}^{n} (y_i - \bar{y})^2 / (n-1)$$

For the Weibull distribution

$$f(x) = (\alpha/\beta) (x/\beta)^{\alpha-1} \exp \left[-(x/\beta)^{\alpha}\right]$$

where α and β are the dispersion and location parameters of x respectively, their maximum likelihood estimators being given by

$$\hat{\alpha} = n \sum_{i=1}^{n} x_i \hat{\alpha} / \left[n \sum_{i=1}^{n} x_i \hat{\alpha} \ln x_i - \left(\sum_{i=1}^{n} \ln x_i \right) / \sum_{i=1}^{n} x_i \hat{\alpha} \right]$$

and

$$n\hat{\beta}^2 = \sum_{i=1}^n x_i^2$$

Estimates of μ and σ and of α and β for the data of Tables 1, 2 and 3 are given in Tables 4, 5 and 6. In these latter, program loading data have been separated from random loading data.

The final steps are taken in Table 7 for the lognormal analysis, and in Table 8 for the Weibull analysis where the results of pooling of variabilities of like groups of test data are presented*. The data have been pooled as follows: For the data treated as log normal

$$\hat{\sigma}^2 = \sum_k \nu_k \, S_k^2 / \sum_k \nu_k \,,$$

and as Weibull (see Appendix)

$$\hat{a} = \sum_{k} n_k / \sum_{k} \left[n_k \left(\sum_{i} x_{ik} \hat{x} \ln x_{ik} \right) / \left(\sum_{i} x_{ik} \hat{x} \right) \right] = \sum_{i} \ln x_{ik}.$$

Considering first the lognormal analysis, Table 7, the additional data confirm the values previously established:

- (1) for gust loading, $\hat{\sigma} = 0.14$
- (2) for manoeuvre loading $\hat{\sigma} = 0.09$
- (3) the ground to air cycle data are too limited to permit generalised conclusions: their inclusion in the earlier categories does not significantly alter the above values, and
- (4) loading sequence (program versus random) has no significant effect upon variability in fatigue life.

The corresponding Weibull analysis gives, for estimates of typical dispersion parameters:

- (1) for gust loading $\hat{a} = 3.9$, and
- (2) for manoeuvre loading $\hat{a} = 6 \cdot 2$.

Rounding these to the nearest whole numbers, $\hat{a} = 4$ for civil flying and $\hat{a} = 6$ for military fighter flying. The former is in agreement with another estimate for civil aircraft¹⁷.

It is of interest to contrast the \hat{a} and $\hat{\sigma}$ estimates for the various groups of data of Tables 4, 5 and 6. This is shown in Figure 1. On average a straight line relationship between $1/\hat{a}$ and $\hat{\sigma}$ fits the data points well, and is given by

$$\hat{a}\hat{\sigma}=0.6$$
.

4. CONCLUSIONS

Analysis of the fatigue lives of over 200 variable amplitude fatigue tests on aluminium alloy aircraft structures indicates that

(1) Variability in fatigue life may be characterised by the following typical values of dispersion parameter:

^{*} In Reference 1, the question of whether scatter in fatigue life varied with life was examined. It was concluded there that the life range of the data was too limited to allow resolution of this point. The additional data included here do not alter this conclusion.

	Data treated as					
Loading Sequence	Log Normal	Weibull â				
Gust Manoeuvre	0·14 0·09	3·9 6·2				

- (2) These values are statistically independent of whether the loading sequence was program or random.
- (3) These results are not significantly altered by the inclusion of the very limited available data from sequences which included the ground to air cycle.

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APPENDIX

Pooling of Weibull Dispersion Estimates

By D. G. Ford

Suppose there are m samples each of size n_k , k = 1 to m, from different Weibull populations with location parameter β_k but a common dispersion parameter, α , then

$$f(x_k) = (\alpha/\beta_k) (x_k/\beta_k)^{\alpha-1} \exp \left[-(x_k/\beta_k)^{\alpha}\right]$$

and the likelihood function is

$$\prod_{k} \left\{ \prod_{i=1}^{n_k} (\alpha/\beta_k) (x_{ik}/\beta_k)^{z-1} \exp\left[-(x_{ik}/\beta_k)^z\right] \right\} = e^L \text{ say.}$$

Then

$$L = \sum_{k} n_k \ln \alpha - \alpha \sum_{k} \ln \beta_k + (\alpha - 1) \sum_{k} \sum_{i} \ln x_{ik} - \sum_{k} \sum_{i} (x_{ik}/\beta_k)^{\alpha}$$
 (1)

Maximum likelihood estimators of α and β_k are found by putting $\partial L/\partial \alpha = 0$ and $\partial L/\partial \beta_k = 0$. For the location parameter

$$\frac{\partial L}{\partial \beta_k} = -\alpha n_k / \beta_k + \alpha \sum_{i=1}^{n_k} x_{ik}^{\alpha} / \beta_k^{\alpha+1} = 0$$

and

the second of th

$$\hat{\beta}_{k}^{\hat{\alpha}} = \sum_{i=1}^{n_{k}} x_{ik}^{\hat{\alpha}} / n_{k}. \tag{2}$$

For the dispersion parameter

$$\frac{\partial L}{\partial \alpha} = \sum_{k} n_k / \alpha - \sum_{k} n_k \ln \beta_k + \sum_{k} \sum_{i=1}^{n_k} \ln x_{ik} - \sum_{k} \sum_{i=1}^{n_k} (x_{ik} / \beta_k)^{\widehat{\alpha}} \ln (x_{ik} / \beta_k) = 0$$

which, after simplifying and substituting for β_k from (2) becomes

$$\hat{\alpha} = \sum_{k} n_k / \sum_{k} n_k \left(\sum_{i=1}^{n_k} x_{ik}^{\hat{\alpha}} \ln x_{ik} \right) / \left(\sum_{i=1}^{n_k} x_{ik}^{\hat{\alpha}} \right) - \sum_{i=1}^{n_k} \ln x_{ik} \right). \tag{3}$$

For its solution Equation (3) requires iterative substitution for \hat{a} : the process is programmed at A.R.L.

TABLE 1
Fatigue Life Data—Gust Loading

Group No.	Reference	Structure	Cycles	Group No.	Reference	Structure	Cycles
1	8 Table VII	Mustang Wing	2418000	5	7 Table 98	Builtup Panel	880000 1008000
		(24S-T)	1485000 1400000	6	7 Table 98	(7075-T6) Builtup Panel (7075-T6)	904000 776000 928000
2	8 Table VII	Mustang Wing	1082000 1016000				
		(24S-T)	820000 801000 1951000	7	9 Table 2	Friendship Wing (7075-T6)	195182 247360
			1994000 1396000	8	8 Table VIII	Mustang Wing (24S-T)	5344000 2751000 1626000 3707000
3	2 Table 3	Commando Wing (24S-T)	8950350 8652005 10322737 7160280 13067511 10919427				4060000 2423000 2340000 3598000 2249000 1390000
4	3 Table V	Commando Wing (24S-T)	531063 722007 584766 733941	9	9 Table 2	Friendship Wing (7075-T6)	208710 146870

TABLE 2
Fatigue Life Data—Manoeuvre Loading

Group No.	Reference	Structure	Cycles	Group No.	Reference	Structure	Cycles
	6 Table 15	Builtup	63300	12	10 Table 4	Piston	108592
	o rubic 15	Structure	62272			Provost	88400
	1	(7075)	85958			Wing	137296
		()	52805			1	132912
					1		141376
2	6 Table 16	Builtup	63128		:		183376
_		Structure	75936			1	176608
}	}	(7075)	68408	1	1	1	144608
			47288				112208
				l		l	150016
3	7 Table 98	Builtup	480000	1		1	177552
		Panel	736000		1	1	178752
	Ì	(7075–T6)	960000		'	1	160864
ļ				1		j	142464
Ì	Ę		}				167744
4	4 Table VI		52170		ļ		148144
ļ		Wing	40348			1 1	108560
Ì			38549				138560
	į						162352
5	4 Table VI		82332				152752
ļ		Wing	77291		!	1 1	153168 130768
}		5	22402	 			155200
6	4 Table V	Fighter	22682		*		154000
ĺ	1	Tail	38629	<u>{</u> {	i	\ \ \\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\	158400
		l	33992	 			190128
_	A Table V	Fighton	113209]		1	166624
7	4 Table V	Fighter Tail	93770	1	1		122080
1		1 411	67309		1		165824
1			0/307	İ		1	158320
8	4 Table V	Fighter	107560	1			166528
\ °	4 Table V	Tail	74536	ll.	Ì	i i	189328
		1 4"	109572				121552
1			10,0,0				175104
9	4 Table V	Fighter	45459	1	1	1	118192
) ′	, rabic v	Tail	50883				177408
1			35256				131600
							190192
10	4 Table V	Fighter	31680				123088
ì		Tail	37020	N	;		214752
			40160		1		245552
}		}		}}]		0.450:
11	4 Table V		55019	13	14 pp. B-88	F86-F	84591
[Tail	64833		n 20	Horizontal	
	į		68430		B-89	Stabilizer	150384
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TABLE 2 [Continued]

Group No.	R	eference	Structure	Cycles	Group No.	R	eference	Structure	Cycles
14	11	Table 3	Piston Provost Wing	177200 142192 124800 122192	22	16	Table VIII	Box Beam (7075-T6)	69660 36720 55080
15	12	Table 3	Piston Provost	125248 130448	23	16	Table VIII	Box Beam (7075-T6)	72756 52116 49794
			Wing	127984 80384 127760 120960 140800	24	13	Table 1 (i)	Piston Provost Wing	129200 155600 127600 148400
				140400 144400 100400	25	8	Table VIII		131200 1147000
16	15	Table 6	Vampire Wing	165600 430760 396940				Wing (24S-T)	912000 720000 924000 907000
,			(2L65)	590782 520472 459952 406552 354576	26	4	Table VI	Fighter Wing	751000 45963 51299 44765
17	15	Table 7	Vampire Wing	452476 383590 379318	27	15	Table VIII	Vampire Wing	466716 366324
		:	Wing (2L65)	379318 406445 414028 370062 400856 380564	28	13	Table I (ii)	(2L65) Piston Provost Wing	140400 104272 145120 114160 132000
18	16	Table VIII	Box Beam (7075-T6)	92777 56540 62190	29	13	Table I (iii)	Piston Provost Wing	145600 191200 215680 150000
19	16	Table VIII	Box Beam (7075–T6)	76140 83700 105840	30	13	Table 1 (iv), (v)	Piston Provost	130640 160000
20	16	Table VIII	Box Beam (7075-T6)	194040 212520			(vi)	Wing	132640 114432 155760 102880
21	16	Table VIII	Box Beam (7075-T6)	261290 132182			1		87360 117600 179264 149440 201344

TABLE 3
Fatigue Life Data—Loading Including Ground to Air Cycles

Group No.	P	Reference	Structure	cture Cycles Group Reference No.		Reference	Structure	Cycles	
1	5	Table 2	Britania Wing Joint	981205 987440 1294976 1082724 1083724	7	7	Table 104 (Fig. 27)	Builtup panel	218881 394309
2	7	Table 103 (Fig. 26)	Builtup panel	638764 765582				,	
3	9	Table 2	Friendship Wing	124121 115706					
4	9	Table 2	Friendship Wing	100707 79726					
5	8	Table VIII	Mustang Wing	485000 1053000 668000 272000 542000 979000 961000 1101000 775000 855000					
6	4	Table VI	Fighter Wing	37575 58338					

TABLE 4
Gust Loading Data—Analysis

Grave	Numbers of load Levels	Loading No. of		Lognor	mal	Weibull		
Group No.	Up/Down	Sequence	Specs.	$ar{x}$ Log Cycles	s	β Cycles	α	
1	3/3	Program	4	6.175	0.158	1761000	3 · 238	
2	3/3	,,	7	6.085	0.165	1455000	3.027	
3	16/16	,,	6	6.985	0.091	10635000	5.641	
4	16/16	,,	4	5.804	0.069	681000	8.954	
5	36/36	,,	2	5.974	0.042	974000	17.668	
6	19/19	,,	3	5.938	0.042	898000	18 · 343	
7	15/15	,,	2	5.342	0.073	233000	10-128	
8	11/11	Random	10	6.436	0.182	3323000	2.754	
9	15/15	,,	2	5 · 243	0.108	191000	6.828	

TABLE 5
Manoeuvre Loading Data—Analysis

Group	No. of L	oad Levels	Landina	No. of	Lognori	nal	Wei	bull
No.	Up/ Down	Lower Bound	Loading Sequence	No. of Specs.	$ar{x}$ Log Cycles	S	β Cycles	α
1	5/1	Constant	Program	4	4.813	0.088	71300	5 · 582
2	5/1	,,	**	4	4 · 798	0.088	68000	7 · 684
3	11/1	!	**	3	5 · 843	0.152	800000	4 · 30
4	4/1	•	. ,,	3	4.636	0.071	46400	7 · 56
5	10/1	,,	••	2	4.902	0.019	81000	37.97
6	4/1	•	•••	3	4 · 491	0.121	34400	6.01
7	5/1	••	••	3	4.951	0.114	99000	5 · 80
8	5/1	,,	**	3	4.981	0.094	104000	8 · 44
9	4/1	•	••	3	4.637	0.082	46600	8-47
10	5/1	•	••	3	4 - 558	0.052	37800	12.86
11	5/1	••	••	3	4 · 796	0.049	65200	14-63
12	6/6	Gust	• • • • • • • • • • • • • • • • • • • •	41	5-180	0.087	167000	5 · 23
13	5/1	Constant	•	3	5.075	0.131	133000	5 · 58
14	6/6	Gust	• ••	4	5 · 146	0.074	151000	6 · 58
15	6/6	•	••	11	5 · 099	0.084	136000	7.04
16	6/6	, ,,	,,	8	5 · 650	0.070	483000	6.53
17	6/6	•••	••	8	5.580	0.036	393000	19.07
18	4/1	Constant	••	3	4.838	0.114	77000	4.72
19	5/1	,,	• • • • • • • • • • • • • • • • • • • •	3	4 · 943	0.074	94200	7 - 48
20	6/1	••	••	2	5 · 308	0.028	208000	26 · 37
21	6/1	•••	• ••	2	5 · 269	0.209	220000	3 · 52
22	6/1	••	,,	3	4.716	0.141	59100	4 · 70
23	5/1	••	,,	3	4.759	0.090	62700	5.93
24	6/6	Gust	••	5	5 · 140	0.039	144000	12.82
25	12/9	Gust	Random	6	5.946	0.073	955000	6.65
26	4/1	Constant	. •••	3	4 674	0.031	48700	17.18
27	6/6	Gust	••	2	5.616	0.074	439000	9.90
28	6/6	••	••	5	5 · 101	0.061	134000	10.18
29	6/6	••	••	. 5	5 · 276	0.111	212000	4.63
30	6/6	**	••	11	5.132	0.107	152000	4.68

TABLE 6

Data Including Ground to Air Cycles—Analysis

Туре			Loading History					Lognormal		Weibull	
Group No.			ad Levels	Loading	l ner	No. of Specs		s	β	α	
		Up/ Down	Lower Bound	Sequence	GA Cycle		Cycles		Cycles		
1	Gust	7/7	Gust	Program	64	5	6.034	0.049	1140000	9 · 228	
2	,,	19/19	,,	**	Complex	2	5.845	0.056	731000	13 - 249	
3	,,	15/15	••	••	12.5	2	5.079	0.022	122000	34 - 177	
4	,,	15/15	,,	Random	12	2	4.952	0.072	94900	10.270	
5	,,	11/11	••	••	≈ 36	10	5 · 854	0.190	857000	3 · 495	
6	Man-		Con-								
	oeuvre	4/1	stant	Program	13	2	4.670	0.135	52200	5 · 454	
7	,,	11/1	••	Random	Complex	2	5.468	0.181	340000	4.076	

TABLE 7
Structural Fatigue Life Variability—Lognormal Analysis

(a) Gust Loading

Loading Spectrum Upper Lower		Loading S	Dooled	
		Program	Random	Pooled
Gust	Gust	0·120 (28, 21) →	.S. → 0·176 (12, 10)	0.141 (40, 31)

(b) Manoeuvre Loading

Loading Spectrum		Loading Sequence		Pooled	
Upper	Lower	Program	Random	rooled	
			N.S.		
Manoeuvre	Gust	0.079 (77, 71)	↔ 0.094 (29, 24)	0.083 (106, 95)	
		1 N.S.	1 N.S.	1 N.S.	
			N.S.		
Manoeuvre	Constant	0 · 103 (53, 35)	↔ 0.031 (3, 2)	0.100 (56, 37)	
			N.S.		
Pooled		0.087 (130, 106	$) \leftrightarrow 0.090 (32, 26)$	0.088 (162, 132	

(c) Loading including Ground Air Cycles

Loading Spectrum		Loading Sequence		Pooled	
Upper	Lower	Ground Air Cycle	Program	Random	Toolea
Gust	Gust	Simple	0·045 (7, 5) 1 N.S. 0·056 (2, 1)	+0·182 (12, 10) Sig	0.147 (21, 16)
Gust	Gust	Complex	0.056 (2, 1)		}
Manoeuvr	e Constant	Simple	0.135 (2, 1) ←		
	! !		N.	S.	0.160 (4, 2)
Manoeuvr	e Constant	Complex		→0·181 (2, 1)	J

Entries show standard deviation of log life, with number of test specimens and degrees of freedom, respectively, in parentheses.

N.S.: F test comparison of variances not significant at 5% level.

Sig.: F test comparison of variances significant at 5°_{\circ} level.

TABLE 8
Structural Fatigue Life Variability—Weibull Analysis

(a) Gust Loading

Loading Spectrum		Loading Sequence		Pooled
Upper	Lower	Program	Random	100104
Gust	Gust	4.708 (28)	2.987 (12)	3.934 (40)

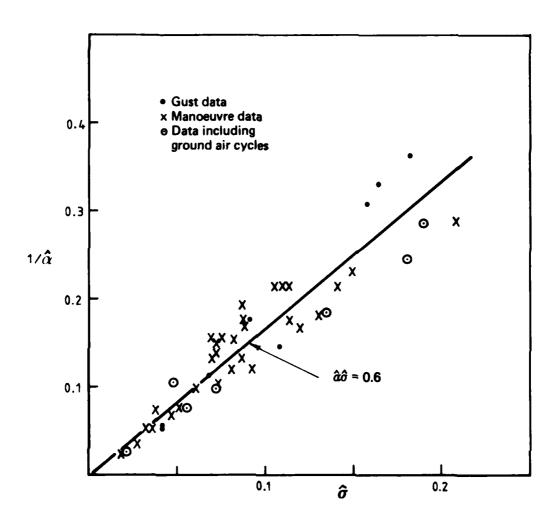
(b) Manoeuvre Loading

Loading Spectrum		Loading Sequence		Pooled
Upper	Lower	Program	Random	Toolog
Manoeuvre	Gust	6.092 (77)	5.620 (29)	5.966 (106
Manoeuvre	Constant	6 · 529 (53)	17 · 185 (3)	6.740 (56)
Pooled		6.233 (130)	5.918 (32)	6 · 171 (162)

(c) Loading Including Ground Air Cycles

Loading Spectrum			Loading Sequence		Pooled
Upper	Lower	Ground Air Cycle	Program	Random	i looku
Gust	Gust	Simple	10.844 (7)	3.893 (12)	5 · 441 (21)
Gust	Gust	Complex	13.25 (2))
Manoeuvre	Constant	Simple	5 · 454 (2)	_	4-665 (4)
Manoeuvre	Constant	Complex		4.076 (2)) . 555 (1)

Entries show dispersion parameter of fatigue life, with number of test specimens in parentheses.



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